

Extreme Value Theory - Maxima

Seminar on Quantitative Risk Management

Franziska Steiner

27th of April 2018



1. Generalized Extreme Value Distribution

2. The Block Maxima Method



Convergence of sums

Let $(X_i)_{i \in \mathbb{N}}$ be a sequence of independent and identically distributed (iid) random variables (rvs) representing financial losses with finite variance.

Further: $S_n = \sum_{i=1}^n X_i$

From the central limit theorem (CLT) :

$$\lim_{n \rightarrow \infty} P \left(\frac{S_n - a_n}{b_n} \leq x \right) = \phi(x), x \in \mathbb{R} \quad (1.1)$$



Generalized Extreme Value Distribution

Definition 1:

The df of the (standard) GEV distribution:

$$H_{\xi}(x) = \begin{cases} \exp(-(1 + \xi x)^{-1/\xi}) & \text{if } \xi \neq 0 \\ \exp(-e^{-x}) & \text{if } \xi = 0 \end{cases}$$

where $1 + \xi x > 0$, ξ is called the shape parameter.

$\xi > 0$: Fréchet distribution

$\xi = 0$: Gumbel distribution

$\xi < 0$: Weibull distribution.



The location parameter $\mu \in \mathbb{R}$ simply shifts the graph left or right on the horizontal axis.

The scale parameter $\sigma > 0$ has the effect to stretch out the graph if $\sigma > 1$ and to squeeze the graph if $\sigma < 1$.



Exponential distribution

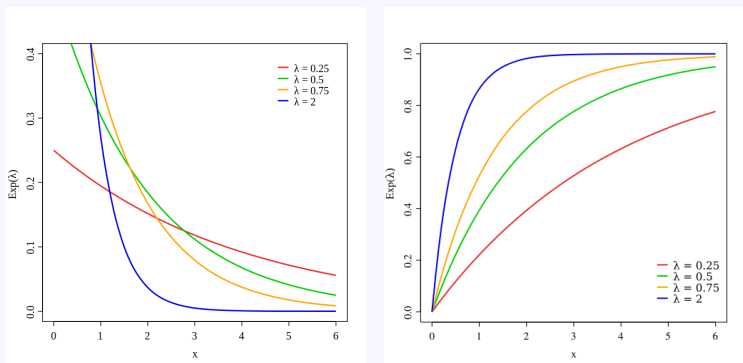


Figure: probability density function(left), cumulative distribution function(right)

source: wikipedia.org



Generalized Extreme Value Distribution

Definition 1:

The df of the (standard) GEV distribution:

$$H_{\xi}(x) = \begin{cases} \exp(-(1 + \xi x)^{-1/\xi}) & \text{if } \xi \neq 0 \\ \exp(-e^{-x}) & \text{if } \xi = 0 \end{cases}$$

where $1 + \xi x > 0$, ξ is called the shape parameter.

$\xi > 0$: Fréchet distribution

$\xi = 0$: Gumbel distribution

$\xi < 0$: Weibull distribution.



Example 1

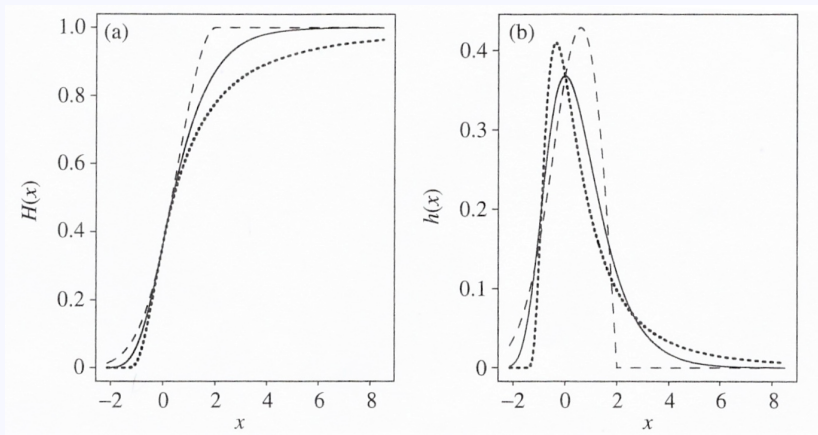


Figure: a) df of the GEV distribution. b) density



Convergence of Maxima

Suppose: block maxima M_n of iid rvs converge in distribution.

We know: $P(M_n \leq x) = F^n(x)$ converges.

\Rightarrow there exist sequences (d_n) and (c_n) with $c_n > 0 \forall n$ such that

$$\lim_{n \rightarrow \infty} P\left(\frac{(M_n - d_n)}{c_n} \leq x\right) = \lim_{n \rightarrow \infty} F^n(c_n x + d_n) = H(x) \quad (1.2)$$

for some non-degenerate df $H(x)$.



Definition 4: maximum domain of attraction

If (1.2) holds for some non-degenerate df H , then F is said to be in the maximum domain of attraction of H , written $F \in MDA(H)$.

Theorem 1: Fisher-Tippett, Gnedenko

If $F \in MDA(H)$ for some non-degenerate df H then H must be a distribution of type H_ξ , i.e. a GEV distribution.



Example 2: The exponential distribution

df: $F(x) = 1 - \exp(-\beta x)$ for $\beta > 0$ and $x \geq 0$.

Choose normalizing sequences $c_n = 1/\beta$ and $d_n = \ln(n/\beta)$ to calculate the limiting distribution of maxima using 1.2.

We get:

$$F^n(c_n x + d_n) = \left(1 - \frac{1}{n} \exp(-x)\right)^n, \quad x \geq -\ln(n)$$

$$\lim_{n \rightarrow \infty} F^n(c_n x + d_n) = \exp(-e^{-x}), \quad x \in \mathbb{R}$$

$$\Rightarrow F \in \text{MDA}(H_0)$$



Convergence of minima

By using the identity $\min(X_1, \dots, X_n) = -\max(-X_1, \dots, -X_n)$ we see that normalized minima of iid samples with df F will convergence in distribution df $\tilde{F}(x) = 1 - F(-x)$.

$$M_n^* = \max(-X_1, \dots, -X_n)$$

Suppose: $\tilde{F} \in MDA(H_\xi)$

$$\text{We get: } \lim_{n \rightarrow \infty} P\left(\frac{M_n^* - d_n}{c_n} \leq x\right) = H_\xi(x)$$

$$\Rightarrow \lim_{n \rightarrow \infty} P\left(\frac{\min(X_1, \dots, X_n) + d_n}{c_n} \leq x\right) = 1 - H_\xi(-x)$$



The Block Maxima Method

Suppose: unknown data from unknown distribution F with $F \in MDA(H_\xi)$.

j -th block of the block maximum: M_{nj} , so our data are M_{n1}, \dots, M_{nm} .

To fit the GEV distribution we use maximum likelihood.

We write $h_{\xi, \mu, \sigma}$ for the density of the GEV distribution.



The log-likelihood:

$$l(\xi, \mu, \sigma; M_{n1}, \dots, M_{nm}) = \sum_{i=1}^m \ln (h_{\xi, \mu, \sigma}(M_{ni}))$$

$$= -m \ln \sigma - \left(1 + \frac{1}{\xi}\right) \sum_{i=1}^m \ln \left(1 + \xi \frac{M_{ni} - \mu}{\sigma}\right) - \sum_{i=1}^m \left(1 + \xi \frac{M_{ni} - \mu}{\sigma}\right)^{-1/\xi}$$

which must be maximized subject to the parameter constraints that $\sigma > 0$ and $1 + \xi (M_{ni} - \mu) / \sigma > 0 \forall i$.



Example 3

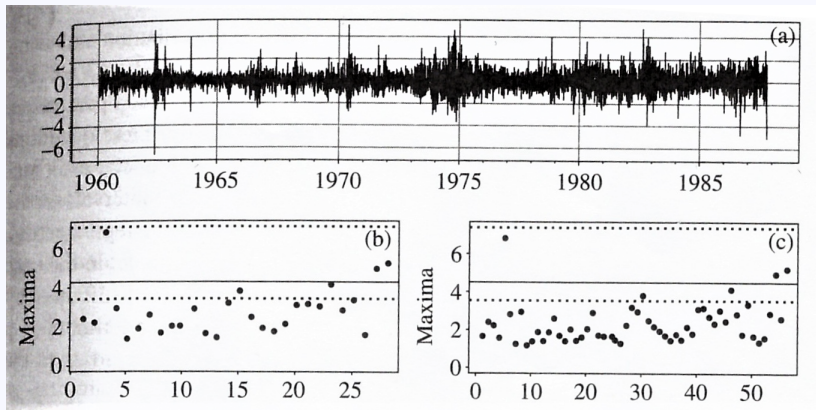


Figure: block maxima analysis of S and P return data



Definition 5 (return level)

Let H denote the df of the true distribution of the n -block maximum. The k n -block return level is $r_{n,k} = q_{1-1/k}(H)$, i.e. the $(1 - 1/k)$ -quantile of H .

Definition 6 (return period)

Let H denote the df of the true distribution of the n -block maximum. The return period of the event $\{M_n > u\}$ is given by $k_{n,u} = 1/\bar{H}(u)$.



Quantitative Risk Management: McNeil, Frey, Embrechts

Uni Muenster Extremwerttheorie, Statistik der
Extremwertverteilungen

