

# On the Monitoring of Structural Changes in Linear Models with Dependent Errors

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# The Sequential Change-Point Problem I

Consider a regression model with MDF errors:

$$y_i = \beta_{i1}x_{i1} + \dots + \beta_{ip}x_{ip} + \varepsilon_i, \quad i = 1, 2, \dots$$

Chu, Stinchcombe & White (1996), historical period:

$$\beta_i = \beta_0, \quad 1 \leq i \leq m.$$

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$$\beta_i = \beta_0, \quad 1 \leq i \leq m.$$

Check parameter stability each time new data arrive:

$$H_0 : \beta_{m+k} = \beta_0 \quad \text{versus} \quad H_A : \beta_{m+k} = \beta_* \neq \beta_0 \quad \forall k \geq k^*.$$

# The Sequential Change-Point Problem II

Horváth, Hušková, Kokoszka & Steinebach (2004) considered linear models with independent errors:

$$y_i = \beta_{i1} \mathbf{1} + \beta_{i2} x_{i2} + \dots + \beta_{ip} x_{ip} + \varepsilon_i, \quad i = 1, 2, \dots$$

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Stopping time of the monitoring procedure:

$$\tau_m = \inf \{k \geq 1 : |Q_m(k)| > c(\alpha) g_m(k)\},$$

such that

$$\lim_{m \rightarrow \infty} \mathbb{P}_{H_0}(\tau_m < \infty) = \alpha$$

and

$$\lim_{m \rightarrow \infty} \mathbb{P}_{H_A}(\tau_m < \infty) = 1.$$

## Detector and Weight Function

Now, we consider a **simple** linear regression with MDF errors:

$$y_i = \beta_i x_i + \varepsilon_i, \quad i = 1, 2, \dots$$

**Residual**-based **CUSUM** detector:  $\hat{\varepsilon}_i = y_i - x_i \hat{\beta}_m$  and

$$Q_m(k) = \sum_{i=1}^{m+k} \left( \hat{\varepsilon}_i - \frac{1}{m} \sum_{i=1}^m \hat{\varepsilon}_i \right), \quad k = 1, 2, \dots$$

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**Horváth et al. (2004)** introduced “early-change” weight function:

$$g_m(k) = m^{1/2} \left( 1 + \frac{k}{m} \right)^{1-\gamma} \left( \frac{k}{m} \right)^{\gamma}, \quad k = 1, 2, \dots,$$

where  $0 \leq \gamma < 1/2$ . Here, we use a “moderate” version:

$$g_m^*(k) = m^{1/2} \left( 1 + \frac{k}{m} \right) \left( \frac{k}{m} \right)^{\gamma}, \quad k = 1, 2, \dots$$

## Assumptions on Regressors

There is a constant  $0 \leq \tau < 1/2$ , such that:

$$\frac{1}{n} \sum_{i=1}^n x_i = o(n^{-\tau}) \quad \text{almost surely}$$

and another constant  $\mathbf{d} > 0$ , such that:

$$\frac{1}{n} \sum_{i=1}^n (x_i^2 - \mathbf{d}) = o(1) \quad \text{almost surely.}$$



## Assumptions on Errors

1. MDF error sequence:

$$\{(\varepsilon_i, \mathcal{F}_i), i \in \mathbb{N}\} \quad \text{with} \quad \mathbb{E}(\varepsilon_i | \mathcal{F}_{i-1}) = 0.$$

2. A moment condition:

$$\sup_{1 \leq i < \infty} \mathbb{E}|\varepsilon_i|^{2+\delta} < \infty \quad \text{for some} \quad 0 < \delta \leq 2.$$

3. Assumption on the conditional variances:

$$\sum_{i=1}^n \{ \mathbb{E}(\varepsilon_i^2 | \mathcal{F}_{i-1}) - \mathbb{E}\varepsilon_i^2 \} = o(\mathbf{a}_n) \quad \text{almost surely,}$$

$$\text{where} \quad \mathbf{a}_n := n^{2/(2+\delta)} (\log n)^{2(1+\delta)/(2+\delta)}.$$

4. Variance of partial sums satisfy

$$\mathbb{E} \left( \sum_{i=1}^n \varepsilon_i \right)^2 = n\sigma^2 + O(\mathbf{a}_n) \quad \text{for some} \quad \sigma^2 > 0.$$

## Asymptotics for Type-1 Error

♣ If  $0 \leq \gamma < \min \left\{ \tau, \frac{1}{2} - \frac{1}{2+\delta} \right\}$ , under  $H_0$ , we have

$$\lim_{m \rightarrow \infty} \mathbb{P} \left( \frac{1}{\sigma} \sup_{1 \leq k < \infty} \frac{|Q_m(k)|}{g_m^*(k)} > c \right) = \mathbb{P} \left( \sup_{0 \leq t \leq 1} (1-t)^\gamma |W(t)| > c \right)$$

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♣ Moreover,

$$\sigma_m^2 := \frac{1}{m} \sum_{i=1}^m \hat{\varepsilon}_i^2 \xrightarrow{\mathbb{P}} \sigma^2 \quad \text{or let} \quad s_m^2(k) := \frac{1}{m+k} \sum_{i=1}^{m+k} \hat{\varepsilon}_i^2$$

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## Example I : NED Regressors

Given sequences  $\{x_i, i \in \mathbb{Z}\}$  and  $\{\eta_i, i \in \mathbb{Z}\}$ , and  $x_i$  is  $\mathcal{F}_{-\infty}^i$ -measurable, where  $\mathcal{F}_{i-k}^i = \sigma\{\eta_j, i-k \leq j \leq i\}$ . Then  $\{x_i, i \in \mathbb{Z}\}$  is called  $\mathcal{L}_p(\nu)$ -near epoch-dependent (NED) on  $\{\eta_i, i \in \mathbb{Z}\}$ , if  $\sup_{i \in \mathbb{Z}} \|x_i\|_p < \infty$  and

$$\sup_{-\infty < i < \infty} \|x_i - \mathbb{E}(x_i | \mathcal{F}_{i-k}^i)\|_p = O(k^{-\nu}) \quad (k \rightarrow \infty),$$

where  $p \geq 1$  and  $\nu > 0$ .

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where  $p \geq 1$  and  $\nu > 0$ .

♣ [improving Ling (2007)] Let  $\{x_i, i \in \mathbb{N}\}$  be a centered  $\mathcal{L}_p(\nu)$ -NED on an independent noise. If  $1 < p < 2$  and  $p/(p-1) \leq q < p(\nu+1)/(p-1)$ , then

$$\frac{1}{n^{\frac{1}{p} + \frac{1}{q}}} \sum_{i=1}^n x_i \rightarrow 0 \quad \text{almost surely.}$$

## Example II : Monitoring AR(1)

Now, consider an autoregressive model, i.e.

$$x_i = \phi x_{i-1} + \eta_i, \quad i = 1, 2, \dots,$$

where  $\{\eta_i, -\infty < i < \infty\}$  is a centered i.i.d.-sequence with:

$$\sigma^2 := \mathbb{E}\eta_1^2, \quad \mathbb{E}|\eta_1|^{2+\delta} < \infty \quad \text{and} \quad -1 < \phi < 1.$$

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Causal representation yields NED-property:

$$\sup_{-\infty < i < \infty} \|x_i - \mathbb{E}(x_i | \mathcal{F}_{i-k}^i)\|_{2+\delta} \ll \phi^k \quad (k \rightarrow \infty).$$

Moreover,  $\{x_i^2 - \sigma^2 / (1 - \phi^2), 1 \leq i < \infty\}$  is  $\mathcal{L}_{1+\delta/2}$ -NED.



## Regressor-weighted CUSUM I

Remember the **simple** linear regression with MDF errors:

$$y_i = \beta_i x_i + \varepsilon_i, \quad i = 1, 2, \dots$$

Following **Hušková & Koubková (2005)** for regressor weights:

$$Q_m(k) = \sum_{i=1}^{m+k} \left( x_i \hat{\varepsilon}_i - \frac{1}{m} \sum_{i=1}^m x_i \hat{\varepsilon}_i \right), \quad k = 1, 2, \dots$$

Now, assume **bounded** regressors and **stationary** MDF errors:

$$\sup_{i \in \mathbb{N}} |x_i| \leq M < \infty, \quad \mathbb{E} \varepsilon_i^2 = \sigma^2 \quad \text{for some } \sigma^2 > 0.$$

## Regressor-weighted CUSUM II

Assume constants  $0 \leq \tau < 1/2$  and  $\mathbf{d} > 0$ , such that

$$\frac{1}{n} \sum_{i=1}^n x_i = o(1) \quad \text{and} \quad \frac{1}{n} \sum_{i=1}^n (x_i^2 - \mathbf{d}) = o(n^{-\tau}).$$

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Assume  $\mathbb{E}|\varepsilon_1|^{2+\delta} < \infty$ , for some  $0 < \delta \leq 2$ , and

$$\sum_{i=1}^n \{ \mathbb{E}(x_i^2 \varepsilon_i^2 | \mathcal{F}_{i-1}) - x_i^2 \sigma^2 \} = o(a_n) \quad \text{almost surely.}$$

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♣ If  $0 \leq \gamma < \delta/(2 + \delta) \leq \tau$ , under  $H_0$ , we have

$$\lim_{m \rightarrow \infty} \mathbb{P} \left( \frac{1}{\sqrt{\mathbf{d}\sigma}} \sup_{1 \leq k < \infty} \frac{|Q_m(k)|}{g_m^*(k)} > c \right) = \mathbb{P} \left( \sup_{0 \leq t \leq 1} (1-t)^\gamma |W(t)| > c \right)$$

## Regressor-weighted CUSUM III

Consider one-time parameter shift alternative, i.e.

$$H_A : \beta_{m+k} = \beta_* \neq \beta_0 \quad \forall k \geq k^*.$$

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$$\lim_{m \rightarrow \infty} \mathbb{P} \left( \frac{1}{\sqrt{\mathbf{d}\sigma}} \sup_{1 \leq k < \infty} \frac{|Q_m(k)|}{g_m^*(k)} < c \right) = 0.$$

♣ Moreover, if  $0 \leq \theta < \delta/(2 + \delta)$ , we have

$$\frac{1}{m} \sum_{i=1}^m x_i^2 \hat{\varepsilon}_i^2 - \mathbf{d}\sigma^2 = o_{\mathbb{P}} \left( m^{-\theta} \right) \quad (m \rightarrow \infty).$$

## Example III : Structural Change in AR(1)

Again, the autoregressive model, i.e.

$$x_i = \phi_i x_{i-1} + \eta_i, \quad i = 1, 2, \dots,$$

where  $\{\eta_i, -\infty < i < \infty\}$  is a centered i.i.d.-sequence with:

$$\mathbb{E}|\eta_1|^{2+\delta} < \infty \quad \text{and} \quad (H_0) \quad -1 < \phi_i = \phi < 1.$$

Hušková & Koubková (2006), one-time shift  $\Delta \neq 0$ , satisfying:

$$-1 < \phi + \Delta < 1 \quad \text{and} \quad \phi + \Delta \neq 0.$$

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$$-1 < \phi + \Delta < 1 \quad \text{and} \quad \phi + \Delta \neq 0.$$

The sequential test presented here is consistent against:

$$H_A : \phi_{m+k} = \phi_* = 0 \neq \phi \quad \forall k \geq k^*.$$



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