

# Backward Invariance Principles in Change-Point Analysis

Alexander Schmitz

University of Cologne

# Change-Point Problem I

Consider the model: location + “white noise”

$$X_i = \beta_i + \varepsilon_i \quad \text{and} \quad \varepsilon_i \sim (\mu, \sigma^2) \quad (i = 1, \dots, n)$$

Test the no-change-in-location null hypothesis

$$H_0 : \beta_1 = \beta_2 = \dots = \beta_n \quad (\beta\text{'s unknown !})$$

against one-time-shift alternative

$$H_A : \text{for some } k \quad \beta_1 = \dots = \beta_k \neq \beta_{k+1} = \dots = \beta_n$$

Based on comparison of the sub-sample means

$$\max_{1 \leq k < n} \left| \frac{1}{k} \sum_{i=1}^k X_i - \frac{1}{n-k} \sum_{i=k+1}^n X_i \right|$$

# Change-Point Problem I

Consider the model: location + “white noise”

$$X_i = \beta_i + \varepsilon_i \quad \text{and} \quad \varepsilon_i \sim (\mu, \sigma^2) \quad (i = 1, \dots, n)$$

Test the no-change-in-location null hypothesis

$$H_0 : \beta_1 = \beta_2 = \dots = \beta_n \quad (\beta\text{'s unknown !})$$

against one-time-shift alternative

$$H_A : \text{for some } k \quad \beta_1 = \dots = \beta_k \neq \beta_{k+1} = \dots = \beta_n$$

Based on comparison of the sub-sample means

$$\max_{1 \leq k < n} \left| \frac{1}{k} \sum_{i=1}^k X_i - \frac{1}{n-k} \sum_{i=k+1}^n X_i \right|$$

Therefore  $H_0 : EX_i = 0 \quad (i = 1, \dots, n)$  and normalized version:

$$\max_{1 \leq k < n} \frac{\left| \frac{1}{k} \sum_{i=1}^k X_i - \frac{1}{n-k} \sum_{i=k+1}^n X_i \right|}{\sigma \sqrt{\frac{1}{k} + \frac{1}{n-k}}}$$

## Change-Point Problem II

Reject the no-change null hypothesis , if

$$\max_{1 \leq k \leq n-1} \frac{n^{-1/2} \left| \sum_{i=1}^k X_i - \frac{k}{n} \sum_{i=1}^n X_i \right|}{\sigma \sqrt{\frac{k}{n} \left(1 - \frac{k}{n}\right)}} = \max_{1 \leq k \leq n-1} \frac{n^{-1/2} |T_n(k)|}{\sigma \sqrt{\frac{k}{n} \left(1 - \frac{k}{n}\right)}}$$

is large.

♣ First approach via FCLT : **constrained versions**

$$\sup_{\varepsilon \leq t \leq 1-\varepsilon} \frac{n^{-1/2} |T_n(\lfloor (n+1)t \rfloor)|}{\sigma \sqrt{t(1-t)}} \xrightarrow{\mathcal{D}} \sup_{\varepsilon \leq t \leq 1-\varepsilon} \frac{|W(t) - tW(1)|}{\sqrt{t(1-t)}}$$

## Change-Point Problem II

Reject the no-change null hypothesis , if

$$\max_{1 \leq k \leq n-1} \frac{n^{-1/2} \left| \sum_{i=1}^k X_i - \frac{k}{n} \sum_{i=1}^n X_i \right|}{\sigma \sqrt{\frac{k}{n} \left(1 - \frac{k}{n}\right)}} = \max_{1 \leq k \leq n-1} \frac{n^{-1/2} |T_n(k)|}{\sigma \sqrt{\frac{k}{n} \left(1 - \frac{k}{n}\right)}}$$

is large.

♣ First approach via FCLT : **constrained versions**

$$\sup_{\varepsilon \leq t \leq 1-\varepsilon} \frac{n^{-1/2} |T_n(\lfloor (n+1)t \rfloor)|}{\sigma \sqrt{t(1-t)}} \xrightarrow{\mathcal{D}} \sup_{\varepsilon \leq t \leq 1-\varepsilon} \frac{|W(t) - tW(1)|}{\sqrt{t(1-t)}}$$

♣ If  $\varepsilon = 0$  and for instance i.i.d.r.v.'s with finite 2nd-moments then

$$(2 \log \log n)^{-1/2} \sup_{\frac{1}{n+1} \leq t \leq \frac{n}{n+1}} \frac{n^{-1/2} |T_n(\lfloor (n+1)t \rfloor)|}{\sigma \sqrt{t(1-t)}} \xrightarrow{P} 1$$

## weighted-sup-norm Donsker I

♣ [following Szyszkowicz(1991) and Horváth(1997)]

Second approach: different weight functions

Let  $0 \leq \tau < 1/2$ . Then

$$\sup_{0 < t < 1} \frac{n^{-1/2} |T_n(\lfloor (n+1)t \rfloor)|}{\sigma(t(1-t))^\tau} \xrightarrow{\mathcal{D}} \sup_{0 < t < 1} \frac{|W(t) - tW(1)|}{(t(1-t))^\tau}$$

## weighted-sup-norm Donsker I

♣ [following Szyszkowicz(1991) and Horváth(1997)]

Second approach: different weight functions

Let  $0 \leq \tau < 1/2$ . Then

$$\sup_{0 < t < 1} \frac{n^{-1/2} |T_n(\lfloor (n+1)t \rfloor)|}{\sigma(t(1-t))^\tau} \xrightarrow{\mathcal{D}} \sup_{0 < t < 1} \frac{|W(t) - tW(1)|}{(t(1-t))^\tau}$$

- “Decomposition” of the tied-down partial sum:

$$T_n(k) = \left(1 - \frac{k}{n}\right) \sum_{i=1}^k X_i - \frac{k}{n} \sum_{i=k+1}^n X_i \quad (1 \leq k < n)$$

## weighted-sup-norm Donsker I

♣ [following Szyszkowicz(1991) and Horváth(1997)]

Second approach: different weight functions

Let  $0 \leq \tau < 1/2$ . Then

$$\sup_{0 < t < 1} \frac{n^{-1/2} |T_n(\lfloor (n+1)t \rfloor)|}{\sigma(t(1-t))^\tau} \xrightarrow{\mathcal{D}} \sup_{0 < t < 1} \frac{|W(t) - tW(1)|}{(t(1-t))^\tau}$$

- “Decomposition” of the tied-down partial sum:

$$T_n(k) = \left(1 - \frac{k}{n}\right) \sum_{i=1}^k X_i - \frac{k}{n} \sum_{i=k+1}^n X_i \quad (1 \leq k < n)$$

- Reversed process is still strongly mixing:  
[Bradley's SIP (1983)]: strict. stat.,  $(2 + \delta)$ -moment and  $\alpha$ -mixing,  
e.g.  $\alpha(n) \ll (\log n)^{-\lambda}$  ( $n \rightarrow \infty$ ) ( $\lambda > 1 + 3/\delta$ )

$$\sum_{i=1}^k X_{-i} - \sigma W(k) = o(\sqrt{k/\log \log k}) \quad a.s. \quad (k \rightarrow \infty)$$



## weighted-sup-norm Donsker II

♣ [convergence-in-probability version]

Let  $0 \leq \tau < 1/2$ . Then there exists a sequence of Brownian bridges on the extended probability space such that:

$$\sup_{0 < t < 1} \frac{|n^{-1/2} T_n(\lfloor (n+1)t \rfloor) - \sigma B_n(t)|}{(t(1-t))^\tau} \xrightarrow{P} 0.$$

## weighted-sup-norm Donsker II

♣ [convergence-in-probability version]

Let  $0 \leq \tau < 1/2$ . Then there exists a sequence of Brownian bridges on the extended probability space such that:

$$\sup_{0 < t < 1} \frac{|n^{-1/2} T_n(\lfloor (n+1)t \rfloor) - \sigma B_n(t)|}{(t(1-t))^\tau} \xrightarrow{P} 0.$$

- $\{W(t) - tW(1), 0 \leq t \leq 1\} \stackrel{D}{=} \{B(t), 0 \leq t \leq 1\}$
- via [Skorokhod's representation of converging laws (1956)] and [Berbee's Extension Lemma (1979)]:

$$\sup_{0 < t < 1} |n^{-1/2} T_n(\lfloor (n+1)t \rfloor) - \sigma B_n(t)| \xrightarrow{P} 0$$

## weighted-sup-norm Donsker II

♣ [convergence-in-probability version]

Let  $0 \leq \tau < 1/2$ . Then there exists a sequence of Brownian bridges on the extended probability space such that:

$$\sup_{0 < t < 1} \frac{|n^{-1/2} T_n(\lfloor (n+1)t \rfloor) - \sigma B_n(t)|}{(t(1-t))^\tau} \xrightarrow{P} 0.$$

- $\{W(t) - tW(1), 0 \leq t \leq 1\} \stackrel{D}{=} \{B(t), 0 \leq t \leq 1\}$
- via [Skorokhod's representation of converging laws (1956)] and [Berbee's Extension Lemma (1979)]:

$$\sup_{0 < t < 1} |n^{-1/2} T_n(\lfloor (n+1)t \rfloor) - \sigma B_n(t)| \xrightarrow{P} 0$$

- via [Bradley's SIP (1983)] and [Berbee's Extension Lemma (1979)]:

$$\max_{1 \leq k \leq n-1} \left| \sum_{i=1}^k X_i - \sigma W_{1n}(k) \right| / \sqrt{k} = O_P(1) \quad (n \rightarrow \infty)$$

$$\max_{1 \leq k \leq n-1} \left| \sum_{i=k+1}^n X_i - \sigma W_{2n}(n-k) \right| / \sqrt{n-k} = O_P(1) \quad (n \rightarrow \infty)$$

## Darling-Erdős Limit Theorem

♣ [following Davis et al.(1986, 1995), Horváth(1993)]

$$A(\log n) \max_{1 \leq k \leq n-1} \frac{n^{-1/2} |T_n(k)|}{\sigma \sqrt{\frac{k}{n} (1 - \frac{k}{n})}} - D(\log n) \xrightarrow{\mathcal{D}} E \vee E' \quad (\spadesuit)$$

where  $E$  and  $E'$  are **independent** i.d.r.v so that

$$P[E \leq y] = P[E' \leq y] = \exp\{\exp(-y)\}$$

and  $A(x) = (2 \log x)^{1/2}$  and  $D(x) = 2 \log x + \frac{1}{2} \log \log x - \frac{1}{2} \log \pi$ .

## Darling-Erdős Limit Theorem

♣ [following Davis et al.(1986, 1995), Horváth(1993)]

$$A(\log n) \max_{1 \leq k \leq n-1} \frac{n^{-1/2} |T_n(k)|}{\sigma \sqrt{\frac{k}{n} (1 - \frac{k}{n})}} - D(\log n) \xrightarrow{\mathcal{D}} E \vee E' \quad (\spadesuit)$$

where  $E$  and  $E'$  are **independent** i.d.r.v so that

$$P[E \leq y] = P[E' \leq y] = \exp\{\exp(-y)\}$$

and  $A(x) = (2 \log x)^{1/2}$  and  $D(x) = 2 \log x + \frac{1}{2} \log \log x - \frac{1}{2} \log \pi$ .

- [Kuelbs & Philipp's SIP (1980)] : strict. stat.,  $2 + \delta$ -moment and  $\alpha$ -mixing, e.g.  $\alpha(n) \ll n^{-(1+\epsilon)(1+2/\delta)}$  ( $n \rightarrow \infty$ )
- **asymptotic independence:**

$$\max_{1 \leq k \leq \frac{n}{\log n}} \frac{\left| \sum_{i=1}^k X_i \right|}{\sqrt{k}} \vee \max_{n - \frac{n}{\log n} \leq k \leq n-1} \frac{\left| \sum_{i=k+1}^n X_i \right|}{\sqrt{n-k}}$$

♣ [Schmitz(2011)] If  $\alpha(n) \ll (\log n)^{-\lambda}$  and  $\beta(n) \downarrow 0$  ( $n \rightarrow \infty$ ) then via [Berbee's Coupling (1979)] we still have  $(\spadesuit)$ .

# Autoregressive Moving Average Models

Given a **weight sequence**  $\{a_k, k \in \mathbb{Z}\}$  we define a **moving average** on i.i.d. noise by

$$X_n = \sum_{k=-\infty}^{\infty} a_k \varepsilon_{n-k}.$$

The ARMA  $(p, q)$   $\{X_n, n \in \mathbb{Z}\}$  is defined as the stationary solution of

$$\phi(B)X_n = \theta(B)\varepsilon_n,$$

where  $\phi(x)$  and  $\theta(x)$  are polynomials of degree  $p$  and  $q \in \{0, 1, \dots\}$  and the constant term of both polynomials is assumed to be one.

# Autoregressive Moving Average Models

Given a **weight sequence**  $\{a_k, k \in \mathbb{Z}\}$  we define a **moving average** on i.i.d. noise by

$$X_n = \sum_{k=-\infty}^{\infty} a_k \varepsilon_{n-k}.$$

The ARMA  $(p, q)$   $\{X_n, n \in \mathbb{Z}\}$  is defined as the stationary solution of

$$\phi(B)X_n = \theta(B)\varepsilon_n,$$

where  $\phi(x)$  and  $\theta(x)$  are polynomials of degree  $p$  and  $q \in \{0, 1, \dots\}$  and the constant term of both polynomials is assumed to be one.

♣ [existence] If  $E \log_+ |\varepsilon_1| < \infty$  and  $\phi(x)$  has no zeros of absolute value one, then there is a stationary ARMA solution. This solution has a moving average representation and is ergodic.

## augmented GARCH

♣ [Duan(1997), Carrasco and Chen (2002)] Let  $\{\eta_k, k \in \mathbb{Z}\}$  be centered i.i.d. random variables. Consider

$$\begin{cases} \varepsilon_k = \sigma_k \eta_k \\ h(\sigma_k) = a(\eta_{k-1}) h(\sigma_{k-1}) + \omega(\eta_{k-1}) \end{cases}$$

where  $h(\cdot)$ ,  $h^{-1}(\cdot)$  and  $a(\cdot)$  and  $\omega(\cdot)$  are nonnegative functions.

$$\sigma_k^2 = (a\eta_{k-1}^2 + b)\sigma_{k-1}^2 + \omega = a\varepsilon_{k-1}^2 + b\sigma_{k-1}^2 + \omega$$



## augmented GARCH

♣ [Duan(1997), Carrasco and Chen (2002)] Let  $\{\eta_k, k \in \mathbb{Z}\}$  be centered i.i.d. random variables. Consider

$$\begin{cases} \varepsilon_k = \sigma_k \eta_k \\ h(\sigma_k) = a(\eta_{k-1}) h(\sigma_{k-1}) + \omega(\eta_{k-1}) \end{cases}$$

where  $h(\cdot)$ ,  $h^{-1}(\cdot)$  and  $a(\cdot)$  and  $\omega(\cdot)$  are nonnegative functions.

$$\sigma_k^2 = (a\eta_{k-1}^2 + b)\sigma_{k-1}^2 + \omega = a\varepsilon_{k-1}^2 + b\sigma_{k-1}^2 + \omega$$

♣ [Aue et al.(2006)]  $E \log_+ a(\eta_1) < 0 \implies P[h(\sigma_k) < \infty] = 1$

$$h(\sigma_k) = \omega(\eta_{k-1}) + \sum_{i=1}^{\infty} a(\eta_{k-1}) \dots a(\eta_{k-i}) \omega(\eta_{k-i-1})$$

## augmented GARCH

♣ [Duan(1997), Carrasco and Chen (2002)] Let  $\{\eta_k, k \in \mathbb{Z}\}$  be centered i.i.d. random variables. Consider

$$\begin{cases} \varepsilon_k = \sigma_k \eta_k \\ h(\sigma_k) = a(\eta_{k-1}) h(\sigma_{k-1}) + \omega(\eta_{k-1}) \end{cases}$$

where  $h(\cdot), h^{-1}(\cdot)$  and  $a(\cdot)$  and  $\omega(\cdot)$  are nonnegative functions.

$$\sigma_k^2 = (a\eta_{k-1}^2 + b)\sigma_{k-1}^2 + \omega = a\varepsilon_{k-1}^2 + b\sigma_{k-1}^2 + \omega$$

♣ [Aue et al.(2006)]  $E \log_+ a(\eta_1) < 0 \implies P[h(\sigma_k) < \infty] = 1$

$$h(\sigma_k) = \omega(\eta_{k-1}) + \sum_{i=1}^{\infty} a(\eta_{k-1}) \dots a(\eta_{k-i}) \omega(\eta_{k-i-1})$$

♣ [Francq & Zakoïan(2006)] augm. GARCH is **geometric ergodic** if

$$dP_\eta = \sum_{i=1}^N p_i \delta_i + (1-p)fd\lambda \quad \text{and} \quad E \log_+ a(\eta_1) < 0.$$

## Linear Process with Dependent Errors

♣ [Francq & Zakoïan(2006)]: log returns behave like AR models with augm. GARCH errors, i.e.

$$\begin{cases} y_k = \phi y_{k-1} + \varepsilon_k, & |\phi| < 1 \\ \varepsilon_k = \sigma_k \eta_k \end{cases}$$

Question: Change-in-the-mean, i.e.  $y_1, \dots, y_k, \Delta + y_{k+1}, \dots, \Delta + y_n$  ?

## Linear Process with Dependent Errors

♣ [Francq & Zakoïan(2006)]: log returns behave like AR models with augm. GARCH errors, i.e.

$$\begin{cases} y_k = \phi y_{k-1} + \varepsilon_k, & |\phi| < 1 \\ \varepsilon_k = \sigma_k \eta_k \end{cases}$$

Question: Change-in-the-mean, i.e.  $y_1, \dots, y_k, \Delta + y_{k+1}, \dots, \Delta + y_n$  ?

♣ [Schmitz(2011) strong approximation results]

- $\{\varepsilon_k\}$  strict. stat.,  $(2 + \delta)$ -moment, **geometric.  $\alpha$ -mixing**
- $\sigma_k = f(\dots, \eta_{k-2}, \eta_{k-1})$  and  $\{\eta_k\}$  centered i.i.d.,

$$\sum_{i=1}^n y_i - \sqrt{\Gamma} W(n) \ll n^{\frac{1}{2}-\kappa} \quad \text{a.s.} \quad (n \rightarrow \infty)$$

## Linear Process with Dependent Errors

♣ [Francq & Zakoïan(2006)]: log returns behave like AR models with augm. GARCH errors, i.e.

$$\begin{cases} y_k = \phi y_{k-1} + \varepsilon_k, & |\phi| < 1 \\ \varepsilon_k = \sigma_k \eta_k \end{cases}$$

Question: Change-in-the-mean, i.e.  $y_1, \dots, y_k, \Delta + y_{k+1}, \dots, \Delta + y_n$  ?

♣ [Schmitz(2011) strong approximation results]

- $\{\varepsilon_k\}$  strict. stat.,  $(2 + \delta)$ -moment, **geometric.  $\alpha$ -mixing**
- $\sigma_k = f(\dots, \eta_{k-2}, \eta_{k-1})$  and  $\{\eta_k\}$  centered i.i.d.,

$$\sum_{i=1}^n y_i - \sqrt{\Gamma} W(n) \ll n^{\frac{1}{2}-\kappa} \quad \text{a.s.} \quad (n \rightarrow \infty)$$

- Mixingale approximation via **truncation**  $\sum_{i=1}^n y_i \rightarrow \sum_{i=1}^n \tilde{y}_i$
- **Big-Block-Small-Block** approach from [Aue et al.(2006)]
- Block **de-coupling** via [Bradley's coupling (1983)]
- [Einmahl's SIP (1989)] for **independent not identic. distributed r.v.**

## “Backward” Invariance Principle

- $\{\varepsilon_k\}$  strict. stat.,  $(2 + \delta)$ -moment, **geometric**.  $\alpha$ -mixing
- $\sigma_k = f(\dots, \eta_{k-2}, \eta_{k-1})$  and  $\{\eta_k\}$  centered i.i.d.,
- $\|\varepsilon_k - E[\varepsilon_k | \eta_{k-m}, \dots, \eta_k]\|_{2+\delta} \leq c_k \psi_m, \quad \sum \psi_m < \infty$

$$\sum_{i=1}^n y_{-i} - \sqrt{\Gamma} W(n) \ll n^{\frac{1}{2}-\kappa} \quad \text{a.s.} \quad (n \rightarrow \infty)$$

## “Backward” Invariance Principle

- $\{\varepsilon_k\}$  strict. stat.,  $(2 + \delta)$ -moment, **geometric.  $\alpha$ -mixing**
- $\sigma_k = f(\dots, \eta_{k-2}, \eta_{k-1})$  and  $\{\eta_k\}$  centered i.i.d.,
- $\|\varepsilon_k - E[\varepsilon_k | \eta_{k-m}, \dots, \eta_k]\|_{2+\delta} \leq c_k \psi_m, \quad \sum \psi_m < \infty$

$$\sum_{i=1}^n y_{-i} - \sqrt{\Gamma} W(n) \ll n^{\frac{1}{2}-\kappa} \quad \text{a.s.} \quad (n \rightarrow \infty)$$

- **reversed “Mixingale” approximation** via truncation

$$\sum_{i=1}^n y_{-i} \rightarrow \sum_{i=1}^n \tilde{y}_{-i}$$

$$\tilde{y}_{-i} = \sum_{\ell=0}^{i^\rho} \phi^\ell \varepsilon_{-i-\ell} \quad \text{for some fixed } \rho, \quad (0 < \rho < 1).$$

- **reversed maximal inequality**

$$\left\| \max_{m \leq j \leq n} \left| \sum_{i=m}^j \tilde{y}_{-i} \right| \right\|_{2+\delta} \leq C_1 \sum_{k=0}^{\infty} \psi_k \left( \sum_{i=1}^{n-m+1} \left( \sum_{\ell=0}^{k \wedge (i+m-1)^\rho} \phi^\ell c_{k-\ell} \right)^2 \right)^{1/2}$$

## Weight Functions I

♣ [Aue et al.(2006)] established SIP for squared augmented GARCH and if  $0 \leq \tau < 1/2$  then

$$(\Gamma n)^{-1/2} \max_{1 \leq k \leq n} \frac{\left| \sum_{i=1}^k \varepsilon_i^2 - \frac{k}{n} \sum_{i=1}^n \varepsilon_i^2 \right|}{\left(\frac{k}{n}\right)^\tau} \xrightarrow{\mathcal{D}} \sup_{0 < t < 1} \frac{|B_t|}{t^\tau}$$

Question: Late change-point, i.e. weight function  $(1-t)^{1/2}$  ?



# Weight Functions I

♣ [Aue et al.(2006)] established SIP for squared augmented GARCH and if  $0 \leq \tau < 1/2$  then

$$(\Gamma n)^{-1/2} \max_{1 \leq k \leq n} \frac{\left| \sum_{i=1}^k \varepsilon_i^2 - \frac{k}{n} \sum_{i=1}^n \varepsilon_i^2 \right|}{\left(\frac{k}{n}\right)^\tau} \xrightarrow{\mathcal{D}} \sup_{0 < t < 1} \frac{|B_t|}{t^\tau}$$

Question: Late change-point, i.e. weight function  $(1-t)^{1/2}$  ?

♣ Answer: Darling-Erdős limit theorem via backward SIP

$$A(\log n) \Gamma^{-1/2} \max_{1 \leq k \leq n-1} \frac{n^{-1/2} \left| \sum_{i=1}^k y_i - \frac{k}{n} \sum_{i=1}^n y_i \right|}{\sqrt{\left(1 - \frac{k}{n}\right)}} \xrightarrow{\mathcal{D}} D^*(\log n) \xrightarrow{\mathcal{D}} E$$

where

$$D^*(x) = 2 \log x + \frac{1}{2} \log \log x - \frac{1}{2} \log(4\pi)$$

Question: Early or late change-point, i.e. weight function  $(t(1-t))^{1/2}$  ?

## Weight Functions II

♣ Answer:

$$A(\log n) \Gamma^{-1/2} \max_{1 \leq k \leq n-1} \frac{n^{-1/2} \left| \sum_{i=1}^k y_i - \frac{k}{n} \sum_{i=1}^n y_i \right|}{\sqrt{\frac{k}{n} \left(1 - \frac{k}{n}\right)}} - D(\log n) \xrightarrow{\mathcal{D}} E \vee E'$$

## Weight Functions II

♣ Answer:

$$A(\log n) \Gamma^{-1/2} \max_{1 \leq k \leq n-1} \frac{n^{-1/2} \left| \sum_{i=1}^k y_i - \frac{k}{n} \sum_{i=1}^n y_i \right|}{\sqrt{\frac{k}{n} \left(1 - \frac{k}{n}\right)}} - D(\log n) \xrightarrow{\mathcal{D}} E \vee E'$$

Asymptotic independence via [Berbee's Extension Lemma (1979)]:





$$\left| \max_{1 \leq k \leq \frac{n}{\log n}} \frac{\left| \sum_{i=1}^k y_i \right|}{\sqrt{k}} - \max_{1 \leq k \leq \frac{n}{\log n}} \frac{|\Gamma^{1/2} W_{1n}(k)|}{\sqrt{k}} \right| = o_P \left( (\log \log n)^{-1/2} \right)$$

the Wiener process  $W_{1n}$  is a function of  $\{\epsilon_1, \dots, \epsilon_{n/\log n}, V\}$ .

$$\left| \max_{n - \frac{n}{\log n} \leq k \leq n-1} \frac{\left| \sum_{i=k+1}^n y_i \right|}{\sqrt{n-k}} - \max_{n - \frac{n}{\log n} \leq k \leq n-1} \frac{|\Gamma^{1/2} W_{2n}(n-k)|}{\sqrt{n-k}} \right| = o_P \left( (\log \log n)^{-1/2} \right)$$

and  $W_{2n}$  is a function of  $\{\epsilon_{n-(n/\log n)+1-(n/\log n)^p}, \dots, \epsilon_n, U\}$ .

# References

-  **Aue, A.; Berkes, I.; Horváth, L.:** Strong approximation for the sums of squares of augmented GARCH sequences. *Bernoulli* 12, 583-608 (2006).
-  **Berbee, H.C.P.:** *Random walks with stationary increments and renewal theory.* Mathematical Centre Tracts 112. Amsterdam: Mathematisch Centrum 1979.
-  **Ling, S.:** Testing for change points in time series models and limiting theorems for NED sequences. *Ann. Statist.* 35, 1213-1237 (2007).
-  **Schmitz, A.:** *Limit Theorems in Change-Point Analysis for Dependent Data.* Dissertation. Universität zu Köln 2011.