

Reversible Markov chains and optimality of symmetrized empirical estimators ^{*}

Priscilla E. Greenwood[†]
University of British Columbia

Wolfgang Wefelmeyer[‡]
Universität-GH Siegen

Abstract

Suppose we want to estimate the expectation of a function of two arguments under the stationary distribution of two successive observations of a reversible Markov chain. Then the usual empirical estimator can be improved by symmetrizing. We show that the symmetrized estimator is efficient. We point out applications to discretely observed continuous-time processes. The proof is based on a result for general Markov chain models which can be used to characterize efficient estimators in any model defined by restrictions on the stationary distribution of a single or two successive observations.

1 Introduction

Suppose we observe X_0, \dots, X_n from an ergodic Markov chain with unknown transition distribution $Q(x, dy)$ and invariant distribution $\pi(dx)$. We want to estimate the expectation of a function $f(x, y)$ under the joint stationary distribution $(\pi \otimes Q)(dx, dy) = \pi(dx)Q(x, dy)$ of two successive observations. Greenwood and Wefelmeyer (1995) show that the *empirical estimator*

$$E_n f = \frac{1}{n} \sum_{i=1}^n f(X_{i-1}, X_i)$$

is asymptotically efficient.

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[†]University of British Columbia, Department of Mathematics, 121-1984 Mathematics Road, Vancouver, B. C., Canada V6T 1Z2.

[‡]Universität - Gesamthochschule Siegen, Fachbereich 6 Mathematik, Hölderlinstr. 3, 57068 Siegen, Germany.

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Suppose now we know the chain to be *reversible*,

$$\pi(dx)Q(x, dy) = \pi(dy)Q(y, dx).$$

This means that the joint stationary distribution of two successive observations is symmetric in the two components. Hence the empirical estimator can be improved by symmetrizing. The *symmetrized empirical estimator* is

$$E_n^s f = \frac{1}{2n} \sum_{i=1}^n (f(X_{i-1}, X_i) + f(X_i, X_{i-1})).$$

We show that this estimator is asymptotically efficient.

For a function $f(x)$ of one argument, the symmetrized empirical estimator is asymptotically equivalent to the ordinary empirical estimator $\frac{1}{n} \sum_{i=1}^n f(X_i)$. Hence our result implies that reversibility carries no information on the expectation of such a function.

The proof is based on a characterization of efficient estimators in any model defined by a restriction on the stationary distribution of a single or two successive observations. Existing efficiency results for ergodic Markov chains have been for two types of model: Fully nonparametric models, and models in which the *transition* distribution is parametrized. Penev (1991) and Bickel (1993) consider estimating πf and Greenwood and Wefelmeyer (1995) consider estimating $\pi \otimes Q f$ for models of the first type, when nothing is known about Q . The main examples of the second type are autoregressive processes, which can be described by parametric or semiparametric models for their transition distributions. For efficient estimation see, e.g., Kreiss (1987), Jeganathan (1995), Koul and Schick (1997) and Drost and Klaassen (1996). Autoregressive processes with non-independent innovations are treated in Wefelmeyer (1996). The techniques for the type of model considered here, where the restriction is on the *stationary* law, are different.

An application of our result is to discretely observed diffusions. Suppose we have a nonparametric model for a diffusion process,

$$dX_t = a(X_t)dt + \sigma(X_t)dW_t.$$

If we observe the process at evenly spaced discrete time points, then the observations form a homogeneous Markov chain. The transition distribution is usually difficult to calculate, and we may be interested in finding the best estimator that does not require this calculation. If a and/or σ vary freely, the corresponding class of invariant distributions includes all distributions with (smooth) positive densities. Since such diffusion processes are reversible, so are the corresponding discretely observed diffusions, and our model consists of reversible Markov chains. A recent review of continuous-time Markov processes is Hansen and Scheinkman (1995). Estimators for parametric models of discretely observed diffusions are studied by Bibby and Sørensen (1995) and Pedersen (1995), and for nonparametric models of discretely observed diffusions by Aït-Sahalia (1996).

In Section 2 we recall a characterization of efficient estimators for the full nonparametric Markov chain model and arbitrary submodels. In Section 3, Theorem 1, we specialize the characterization to submodels defined by restrictions on the stationary distribution of a single or two successive observations. We also describe briefly an application to models defined by restricting the stationary distribution of a single observation to a parametric family. Such models arise when a diffusion is discretely observed and the diffusion model is parametric. Details will be given in Kessler and Wefelmeyer (1997). In Section 4, Theorem 2, we prove efficiency of the symmetrized empirical estimator by applying Theorem 1 to the submodel given by all reversible Markov chains. Section 5 contains the proofs of the lemmas.

2 Full model and submodel

In this section we recall a characterization of efficient estimators for arbitrary smooth functionals of the transition distribution in general Markov chain models. The characterization is based on a nonparametric version of Hájek's (1970) convolution theorem. It requires the model to be locally asymptotically normal.

We begin with the full nonparametric model, in which no restrictions are made on the transition distribution. Let X_0, \dots, X_n be observations from a Markov chain with values in some measurable space. We will introduce a local parameter space at a fixed transition distribution $Q(x, dy)$ such that the chain satisfies the two assumptions below.

Assumption 1. The chain is positive Harris recurrent.

Let $\pi(dx)$ denote the invariant distribution. For a suitably integrable function $f_1(x)$ of *one* argument write

$$(Qf_1)(x) = \int Q(x, dy)f_1(y), \quad \pi f_1 = \int \pi(dx)f_1(x).$$

For a function $f(x, y)$ of *two* arguments we will write

$$(Qf)(x) = \int Q(x, dy)f(x, y), \quad Q^j f = Q^{j-1}Qf, \quad \pi \otimes Qf = \iint \pi(dx)Q(x, dy)f(x, y).$$

Assumption 2. The chain has the ergodicity property

$$\sup\{\pi(Q^j f_1 - \pi f_1)^2 : \pi f_1^2 \leq 1\} \rightarrow 0 \quad \text{for } j \rightarrow \infty.$$

The space

$$H = \{h \in L_2(\pi \otimes Q) : Qh = 0\}$$

is a *local parameter space* in the sense that for each $h \in H$ we can choose a sequence Q_{nh} which is *Hellinger differentiable* with *derivative* h ,

$$\int Q(x, dy) \left((dQ_{nh}/dQ)(x, y)^{1/2} - 1 - n^{-1/2} \frac{1}{2} h(x, y) \right)^2 \leq n^{-1} r_n(x), \quad (2.1)$$

with r_n decreasing to 0 pointwise and π -integrable for large n . This version of differentiability is due to Höpfner, Jacod and Ladelli (1990).

Remark. In the full model we can take

$$Q_{nh}(x, dy) = Q(x, dy)(1 + n^{-1/2} h_n(x, y)) \quad (2.2)$$

with

$$h_n = \bar{h}_n - Q\bar{h}_n \quad \text{and} \quad \bar{h}_n = h1_{(|h| \leq n^{1/8})}. \quad (2.3)$$

□

Write P_n and P_{nh} for the joint distribution of X_0, \dots, X_n under Q and Q_{nh} , respectively. As in Höpfner (1993) we have a nonparametric version of *local asymptotic normality*,

$$\log dP_{nh}/dP_n = n^{-1/2} \sum_{i=1}^n h(X_{i-1}, X_i) - \frac{1}{2} \pi \otimes Q h^2 + o_{P_n}(1)$$

and

$$n^{-1/2} \sum_{i=1}^n h(X_{i-1}, X_i) \Rightarrow N_h,$$

where N_h is normal with mean 0 and variance $\pi \otimes Q h^2$. Local asymptotic normality induces an inner product $\pi \otimes Q h h'$ on H . Efficient estimators for a functional $t(Q)$ are characterized in terms of the gradient of the functional with respect to this inner product. A functional $t(Q)$ of the transition distribution is *differentiable*, with *gradient* $g \in H$, if

$$n^{1/2}(t(Q_{nh}) - t(Q)) \rightarrow \pi \otimes Q h \cdot g \quad \text{for } h \in H. \quad (2.4)$$

The gradient is uniquely determined.

We can now characterize regular and efficient estimators of $t(Q)$. An estimator T_n is called *regular* for $t(Q)$ in the full model, with *limit* L , if

$$n^{1/2}(T_n - t(Q_{nh})) \Rightarrow L \quad \text{under } P_{nh} \text{ for } h \in H. \quad (2.5)$$

A nonparametric version of Hájek's (1970) convolution theorem says that

$$L = N_g + M,$$

where M is independent of N_g and N_g is normal with mean 0 and variance $\pi \otimes Q g^2$. This justifies calling T_n *efficient* if $L = N_g$. It is a well-known consequence of the convolution theorem that an estimator T_n is regular and efficient for $t(Q)$ if and only if

$$n^{1/2}(T_n - t(Q)) = n^{-1/2} \sum_{i=1}^n g(X_{i-1}, X_i) + o_{P_n}(1). \quad (2.6)$$

We now characterize efficient estimators for a *submodel* of the full nonparametric model above. Fix a transition distribution $Q(x, dy)$ which fulfills Assumptions 1 and 2. The local parameter space of the submodel is a subset H_0 of H which we take to be closed and linear. A functional $t(Q)$ is *differentiable* on the submodel, with *gradient* $g \in H$, if (2.4) holds for $h \in H_0$. Now the gradient is not uniquely determined. A special role is played by the *canonical* gradient, the projection of an arbitrary gradient into H_0 .

An estimator T_n is called *regular* for $t(Q)$ in the submodel, with *limit* L , if (2.5) holds for $h \in H_0$. The convolution theorem then says that $L = N_{g_0} + M$, where M and N_{g_0} are independent. Hence T_n is regular and efficient for $t(Q)$ in the submodel if and only if (2.6) holds with g replaced by the canonical gradient g_0 .

Call an estimator *asymptotically linear* for $t(Q)$ with *influence function* $k \in H$ if

$$n^{1/2}(T_n - t(Q)) = n^{-1/2} \sum_{i=1}^n k(X_{i-1}, X_i) + o_{P_n}(1). \quad (2.7)$$

In these terms, an estimator is regular and efficient in the submodel if and only if it is asymptotically linear with influence function equal to the canonical gradient,

$$n^{1/2}(T_n - t(Q)) = n^{-1/2} \sum_{i=1}^n g_0(X_{i-1}, X_i) + o_{P_n}(1). \quad (2.8)$$

Moreover, an asymptotically linear estimator is regular if and only if its influence function is a gradient. A convenient reference for the convolution theorem and the characterizations of regular and efficient estimators is Greenwood and Wefelmeyer (1990).

3 Restricting the invariant law

In this section we consider a submodel described by a restriction on $\pi \otimes Q$. We have recalled in Section 2 that a regular and efficient estimator for the functional $t(Q)$ is characterized as an asymptotically linear estimator with influence function equal to the canonical gradient of $t(Q)$. We calculate the canonical gradient in two steps. First we describe the local parameter space H_0 in terms of the restriction on $\pi \otimes Q$. Then we calculate the projection of an arbitrary gradient into H_0 .

Consider a submodel as in Section 2. Fix a transition distribution $Q(x, dy)$ in the submodel which fulfills Assumptions 1 and 2. Let H_0 denote a local parameter space

at Q . Suppose the submodel is described by a restriction on $\pi \otimes Q$. How can H_0 be described in terms of this restriction? To answer this, we perturb $\pi \otimes Q$ and interpret the restriction on the perturbation as a restriction on H . For smooth models it will be possible to choose, for each $h \in H_0$, a sequence Q_{nh} in the submodel which is Hellinger differentiable in the sense of (2.1) and has the following two additional properties. Each Q_{nh} has an invariant distribution π_{nh} , and $\pi_{nh} \otimes Q_{nh}$ admits a perturbation expansion: For $f \in L_2(\pi \otimes Q)$,

$$n^{1/2}(\pi_{nh} \otimes Q_{nh} f - \pi \otimes Q f) \rightarrow \pi \otimes Q h \cdot A f \quad \text{for } h \in H_0. \quad (3.1)$$

See the Remark below. Here the operator A is a kernel on $L_2(\pi \otimes Q)$,

$$A = I_2 - \pi \otimes Q + (R - L)V, \quad (3.2)$$

where I_2 is the identity on the space $L_2(\pi \otimes Q)$ of functions of *two* arguments, R and L map a function $v \in L_2(\pi)$ of one argument into functions of two arguments,

$$(Rv)(x, y) = v(y), \quad (Lv)(x, y) = v(x),$$

and V is the kernel

$$V = \sum_{j=1}^{\infty} (Q^j - \pi \otimes Q).$$

Writing

$$(Af)(x, y) = f(x, y) - (Qf)(x) + \sum_{j=1}^{\infty} \left((Q^j f)(y) - (Q^{j+1} f)(x) \right),$$

we see that A maps $L_2(\pi \otimes Q)$ into the local parameter space H . Let $L_{2,0}(\pi \otimes Q)$ denote the functions in $L_2(\pi \otimes Q)$ with expectation 0 under $\pi \otimes Q$. Since A is the identity on H , it maps $L_{2,0}(\pi \otimes Q)$ *onto* H . Note that V is similar to the potential $\sum_{j=0}^{\infty} Q^j$ and has the similar property

$$(I + V)Q = \pi \otimes Q + V. \quad (3.3)$$

Remark. In the full nonparametric model, Q_{nh} has the additional properties described above if it is defined as in (2.2) and (2.3). Then

$$\pi(Q_{nh}f - Qf)^2 = n^{-1}\pi(Qh_n f)^2 \leq 2n^{-3/4}\pi f^2.$$

Hence Assumption 2 holds and Q_{nh} has an invariant distribution π_{nh} . Further, the perturbation expansion (3.1) follows from Lemma 2 in Greenwood and Wefelmeyer (1997), which is based on a version of a perturbation expansion of $\pi_{nh}f$ of Kartashov (1985a, 1985b). \square

Let B denote the adjoint of A . Then (3.1) reads

$$n^{1/2}(\pi_{nh} \otimes Q_{nh} f - \pi \otimes Q f) \rightarrow \pi \otimes Q B h \cdot f \quad \text{for } h \in H_0. \quad (3.4)$$

This means that the $\pi \otimes Q$ -density of $\pi_{nh} \otimes Q_{nh}$ is approximately $1 + n^{-1/2} B h$. Now we see how to translate a restriction on $\pi_{nh} \otimes Q_{nh}$ into a restriction on h through a condition on $B h$. In the examples we have in mind, and in particular in Section 4,

$$H_0 = \{h \in H : B h \in F\} \quad (3.5)$$

for some closed linear subspace F of $L_2(\pi \otimes Q)$.

Once we have described the local parameter space H_0 , we must calculate the canonical gradient of the functional $t(Q)$ we want to estimate. According to Section 2, the canonical gradient g_0 is the projection of an arbitrary gradient $g \in H$ into H_0 . It is therefore characterized by the properties that $g_0 \in H_0$ and $g - g_0$ is orthogonal to H_0 . The operator A maps $L_{2,0}(\pi \otimes Q)$ onto H , so we can write $g = A f$ and $g_0 = A f_0$ for some $f, f_0 \in L_{2,0}(\pi \otimes Q)$. Hence the canonical gradient is characterized as the function $A f_0$ with $f_0 \in L_{2,0}(\pi \otimes Q)$ such that $A f_0 \in H_0$ and $A f - A f_0$ is orthogonal to H_0 . The condition $A f_0 \in H_0$ can, in turn, according to (3.5), be written as $B A f_0 \in F$.

To describe the canonical gradient more usefully, it remains to calculate B and $B A$. The adjoint B will be written in terms of the *reversed* chain, with transition distribution $\overline{Q}(y, dx)$ defined by

$$\pi(dx)Q(x, dy) = \pi(dy)\overline{Q}(y, dx). \quad (3.6)$$

One checks that the transition distribution \overline{Q} also has invariant distribution π .

For a function $f(x, y)$ of two arguments we will follow the convention that the transition distribution of the reversed chain acts on f from right to left, i.e., on the first argument of f ,

$$(\overline{Q}f)(y) = \int \overline{Q}(y, dx) f(x, y).$$

Similarly,

$$\pi \otimes \overline{Q} f = \iint \pi(dy)\overline{Q}(y, dx) f(x, y).$$

The kernel corresponding to V for the reversed chain is

$$\overline{V} = \sum_{j=1}^{\infty} (\overline{Q}^j - \pi \otimes \overline{Q}).$$

From (3.6) we obtain

$$\pi(dx)Q^j(x, dy) = \pi(dy)\overline{Q}^j(y, dx) \quad \text{for } j = 1, 2, \dots \quad (3.7)$$

and therefore

$$\pi(dx)V(x, dy) = \pi(dy)\overline{V}(y, dx). \quad (3.8)$$

The property corresponding to (3.3) is

$$(I + \bar{V})\bar{Q} = \pi \otimes \bar{Q} + \bar{V}. \quad (3.9)$$

Let B denote the kernel on $L_2(\pi \otimes Q)$,

$$B = I_2 - \pi \otimes Q + L(I + \bar{V})(\bar{Q} - Q).$$

By (3.9), on H this kernel reduces to

$$B = I_2 + L(I + \bar{V})\bar{Q} = I_2 + L\bar{V}. \quad (3.10)$$

Lemma 1. *The adjoint of A on $L_2(\pi \otimes Q)$ is B .*

Lemma 2. *We have $BA = I_2 - \pi \otimes Q + RV + L\bar{V}$ on $L_2(\pi \otimes Q)$.*

Lemmas 1 and 2 are proved in Section 4. We arrive at our description of the canonical gradient.

Theorem 1. *Suppose the local parameter space is of the form $H_0 = \{h \in H : Bh \in F\}$ for some closed linear subspace F of $L_2(\pi \otimes Q)$. Let $t(Q)$ be differentiable, and let $g = Af$ be a gradient of $t(Q)$, with $f \in L_{2,0}(\pi \otimes Q)$.*

Then the canonical gradient g_0 is characterized as the function of the form $g_0 = Af_0$ with $f_0 \in L_{2,0}(\pi \otimes Q)$ fulfilling

$$(BAf_0 =) f_0 - \pi \otimes Q f_0 + RV f_0 + L\bar{V} f_0 \in F$$

and

$$\pi \otimes Q A(f - f_0) \cdot h = 0 \quad \text{for } h \in H_0.$$

Example. An application of Theorem 1 is to the submodel of all Markov chains with invariant distribution in some parametric family π_ϑ , $\vartheta \in \Theta$. Such a model arises when a diffusion is discretely observed and the diffusion model is parametric. For simplicity, let Θ be an open subset of the real line. Fix a transition distribution $Q(x, dy)$ with invariant distribution π_ϑ . Let H_0 denote the local parameter space at Q . The perturbation expansion (3.1) of $\pi_{nh} \otimes Q_{nh}$ implies a perturbation expansion of π_{nh} : For $f_1 \in L_2(\pi_\vartheta)$,

$$n^{1/2}(\pi_{nh} f_1 - \pi_\vartheta f_1) \rightarrow \pi_\vartheta h \cdot ARf_1 = \pi_\vartheta Bh \cdot Rf_1 \quad \text{for } h \in H_0.$$

On the other hand, π_{nh} is in the parametric family and therefore approximately of the form $\pi_{\vartheta+n^{-1/2}u}$ for some u . Under a suitable smoothness assumption on the family, a Taylor expansion gives for $f_1 \in L_2(\pi_\vartheta)$,

$$n^{1/2}(\pi_{\vartheta+n^{-1/2}u} f_1 - \pi_\vartheta f_1) \rightarrow u \pi_\vartheta \ell'_\vartheta \cdot f_1 = u \pi_\vartheta \otimes Q R \ell'_\vartheta \cdot Rf_1,$$

where $\ell'_\vartheta = (\partial/\partial\tau)(d\pi_\tau/d\pi_\vartheta)|_{\tau=\vartheta}$. Hence

$$H_0 = \{h \in H : Bh = uR\ell'_\vartheta \text{ for some } u\}.$$

We can now apply Theorem 1 with $F = \{uR\ell'_\vartheta : u \in \mathbf{R}\}$. For details see Kessler and Wefelmeyer (1997). \square

The following observations will be applied in Section 4. Since Q and \overline{Q} have invariant distribution π , so has

$$Q^s = \frac{1}{2}(Q + \overline{Q}).$$

We have

$$\pi(dx)Q(x, dy) = \pi(dy)\overline{Q}(y, dx) = \pi(dx)\overline{\overline{Q}}(x, dy).$$

Hence $\overline{\overline{Q}} = Q$, and

$$\begin{aligned} \pi(dx) \left(Q(x, dy) + \overline{Q}(x, dy) \right) &= \pi(dy) \left(\overline{Q}(y, dx) + \overline{\overline{Q}}(y, dx) \right) \\ &= \pi(dy) \left(Q(y, dx) + \overline{Q}(y, dx) \right). \end{aligned}$$

This means that Q^s is reversible. The transition distribution Q^s is called the *additive reversibilization* of Q by Fill (1991) who uses reversibilizations to obtain bounds on the rate of convergence of Markov chains to stationarity.

4 Reversible Markov chains

Suppose that X_0, \dots, X_n come from a *reversible* Markov chain. This means

$$\pi(dx)Q(x, dy) = \pi(dy)Q(y, dx) \tag{4.1}$$

and therefore $\overline{Q} = Q$ and $\pi \otimes Q$ is symmetric in the two components. We must keep in mind that \overline{Q} still acts differently from Q on functions of two arguments,

$$(\overline{Q}f)(x) = \int \overline{Q}(x, dy)f(y, x) = (Q\overline{f})(x),$$

where $\overline{f}(x, y) = f(y, x)$.

Fix a reversible transition distribution $Q(x, dy)$ which fulfills Assumptions 1 and 2. We have (3.4),

$$n^{1/2}(\pi_{nh} \otimes Q_{nh} f - \pi \otimes Q f) \rightarrow \pi \otimes Q Bh \cdot f \quad \text{for } h \in H. \tag{4.2}$$

If Q_{nh} as well as Q is reversible, then $\pi_{nh} \otimes Q_{nh}$ and $\pi \otimes Q$ are symmetric. Therefore, B must also be symmetric. This suggests that the local parameter space for the model of reversible Markov chains is

$$H_0 = \{h \in H : Bh \text{ symmetric}\}.$$

Indeed, if $h \in H_0$, then $\pi_{nh} \otimes Q_{nh}$ is symmetric up to $o(n^{-1/2})$ by (4.2). It remains to check that for each $h \in H_0$ we can find a Hellinger differentiable sequence Q_{nh}^s such that (4.2) holds and $\pi_{nh} \otimes Q_{nh}^s$ is *exactly* symmetric, so that Q_{nh}^s is exactly reversible. This is the content of Lemma 3.

Lemma 3. *For each $h \in H_0$ there exists a sequence of transition distributions Q_{nh}^s with invariant distribution π_{nh} such that Q_{nh}^s is reversible (4.1) and Hellinger differentiable (2.1) with derivative h , and such that the perturbation expansion (3.1) holds for $\pi_{nh} \otimes Q_{nh}^s$.*

The proof is in Section 4.

Recall that the *empirical estimator* for $\pi \otimes Q f$,

$$E_n f = \frac{1}{n} \sum_{i=1}^n f(X_{i-1}, X_i),$$

is efficient in the full nonparametric model. To prove efficiency of the *symmetrized* empirical estimator $E_n^s f = E_n f_s$ with $f_s = \frac{1}{2}(f + \bar{f})$ and $\bar{f}(x, y) = f(y, x)$ in the model given by all *reversible* Markov chains, we apply Theorem 1.

Theorem 2. *For $f \in L_2(\pi \otimes Q)$, the estimator*

$$E_n f_s = \frac{1}{2n} \sum_{i=1}^n (f(X_{i-1}, X_i) + f(X_i, X_{i-1}))$$

is regular and efficient for $\pi \otimes Q f$ in the model given by all reversible Markov chains.

Proof. By Lemma 3 we have $H_0 = \{h \in H : Bh \text{ symmetric}\}$. Hence the assumption on H_0 of Theorem 1 holds with

$$F = \{f \in L_2(\pi \otimes Q) : f \text{ symmetric}\}.$$

We want to show that Af_s is the canonical gradient of the functional $t(Q) = \pi \otimes Q f$. By the perturbation expansion (3.1), $g = Af$ is a gradient of $\pi \otimes Q f$. It suffices to check that the two conditions on f_0 in Theorem 1 hold for $f_0 = f_s$. For the first condition, we observe that $RVf_s + L\bar{V}f_s$ is symmetric since f_s is symmetric. The second condition holds since by Lemma 1 we have for all $h \in H_0$,

$$\pi \otimes Q (Af - Af_s) \cdot h = \frac{1}{2} \pi \otimes Q A(f - \bar{f}) \cdot h = \frac{1}{2} \pi \otimes Q (f - \bar{f}) \cdot Bh = 0.$$

The last equality holds because Bh is symmetric for $h \in H_0$ and $\pi \otimes Q$ is symmetric, and therefore

$$\pi \otimes Q f \cdot Bh = \pi \otimes Q \bar{f} \cdot Bh.$$

Hence Af_s is the canonical gradient, and the characterization (2.8) says that an estimator for $\pi \otimes Q f$ is regular and efficient if it is asymptotically linear with influence function Af_s .

We have

$$n^{1/2}(E_n f_s - \pi \otimes Q f) = n^{-1/2} \sum_{i=1}^n (f_s(X_{i-1}, X_i) - \pi \otimes Q f).$$

By the martingale approximation of Gordin and Lifšic (1978) we have for $f \in L_{2,0}(\pi \otimes Q)$,

$$n^{-1/2} \sum_{i=1}^n f(X_{i-1}, X_i) = n^{-1/2} \sum_{i=1}^n (Af)(X_{i-1}, X_i) + o_{P_n}(1).$$

This approximation, applied to $f = f_s - \pi \otimes Q f$, shows that $E_n f_s$ has the influence function Af_s . \square

5 Proofs of the lemmas

Proof of Lemma 1. We must prove that for $e, f \in L_2(\pi \otimes Q)$,

$$\pi \otimes Q e \cdot Af = \pi \otimes Q B e \cdot f.$$

Clearly, I_2 is its own adjoint. We prove that $L(\bar{Q} - Q)$ is the adjoint of $(R - L)(Q - \pi \otimes Q)$ and $L(\bar{Q}^{j-1} - \pi \otimes \bar{Q})(\bar{Q} - Q)$ is the adjoint of $(R - L)(Q^j - \pi \otimes Q)$ for $j = 2, 3, \dots$. We can omit the centering by $\pi \otimes Q$ because $(R - L)\pi \otimes Q = 0$. We can omit the centering by $\pi \otimes \bar{Q}$ because also

$$\begin{aligned} & \pi \otimes \bar{Q}(\bar{Q} - Q)e \\ &= \int \pi(dx) \bar{Q}(x, dy) \left(\int \bar{Q}(y, dz) e(y, z) - \int Q(y, dz) e(y, z) \right) \\ &= \int \pi(dy) \left(\int Q(y, dz) e(y, z) - \int \bar{Q}(y, dz) e(y, z) \right) \\ &= 0. \end{aligned}$$

Now

$$\begin{aligned} & \pi \otimes Q e \cdot (R - L)Qf \\ &= \iint \pi(dx) Q(x, dy) e(x, y) \left(\int Q(y, dz) f(y, z) - \int Q(x, dz) f(x, z) \right) \\ &= \int \pi(dy) \int \bar{Q}(y, dx) e(x, y) \int Q(y, dz) f(y, z) \\ &\quad - \int \pi(dx) \int Q(x, dy) e(x, y) \int Q(x, dz) f(x, z) \\ &= \iint \pi(dx) Q(x, dy) (\bar{Q}e)(x) f(x, y) \\ &\quad - \iint \pi(dx) Q(x, dy) (Qe)(x) f(x, y) \\ &= \pi \otimes Q L(\bar{Q} - Q)e \cdot f \end{aligned}$$

and, for $j = 2, 3, \dots$,

$$\begin{aligned}
& \pi \otimes Q e \cdot (R - L)Q^j f \\
&= \iint \pi(dx)Q(x, dy)e(x, y) \left(\int Q^{j-1}(y, dz) \int Q(z, dw)f(z, w) \right. \\
&\quad \left. - \int Q^{j-1}(x, dz) \int Q(z, dw)f(z, w) \right) \\
&= \int \pi(dy) \int \bar{Q}(y, dx)e(x, y) \int Q^{j-1}(y, dz) \int Q(z, dw)f(z, w) \\
&\quad - \int \pi(dz) \int \bar{Q}^{j-1}(z, dx) \int Q(x, dy)e(x, y) \int Q(z, dw)f(z, w) \\
&= \int \pi(dz) \int \bar{Q}^{j-1}(z, dy) \int \bar{Q}(y, dx)e(x, y) \int Q(z, dw)f(z, w) \\
&\quad - \int \pi(dz) \int \bar{Q}^{j-1}(z, dx) \int Q(x, dy)e(x, y) \int Q(z, dw)f(z, w) \\
&= \pi \otimes Q L \bar{Q}^{j-1} (\bar{Q} - Q) e \cdot f.
\end{aligned}$$

□

Proof of Lemma 2. Note that $QR = Q$ and $QL = I_1$ on $L_2(\pi)$, where I_1 is the identity on $L_2(\pi)$. Hence $Q^j R = Q^j$ and $Q^j L = Q^{j-1}$ on $L_2(\pi)$ for $j = 2, 3, \dots$, and therefore

$$VL = \sum_{j=1}^{\infty} (Q^j L - \pi) = I_1 - \pi + VR.$$

For \bar{V} we have the corresponding relation

$$\bar{V}R = I_1 - \pi + \bar{V}L.$$

Hence

$$\bar{V}(R - L) = I_1 - \pi,$$

so that

$$\bar{V}A = \bar{V} + \bar{V}(R - L)V = \bar{V} + V$$

and with (3.10),

$$\begin{aligned}
BA &= (I_2 + L\bar{V})A \\
&= I_2 - \pi \otimes Q + (R - L)V + L\bar{V} + LV \\
&= I_2 - \pi \otimes Q + RV + L\bar{V}.
\end{aligned}$$

□

Proof of Lemma 3. Let Q_{nh} be the sequence $Q(x, dy)(1 + n^{-1/2}h_n(x, y))$ defined in (2.2) and (2.3). By the Remark in Section 3, each transition distribution Q_{nh} has an

invariant distribution π_{nh} , and $\pi_{nh} \otimes Q_{nh}$ admits the perturbation expansion (3.1). The transition distribution \overline{Q}_{nh} of the reversed chain is defined by (3.6),

$$\pi_{nh}(dx)Q_{nh}(x, dy) = \pi_{nh}(dy)\overline{Q}_{nh}(y, dx). \quad (5.1)$$

Let $Q_{nh}^s = \frac{1}{2}(Q_{nh} + \overline{Q}_{nh})$ be the additive reversibilization of Q_{nh} . We have seen at the end of Section 3 that Q_{nh}^s has invariant distribution π_{nh} and is reversible.

By (5.1), (3.1), Lemma 1, reversibility of Q , and symmetry of Bh ,

$$\begin{aligned} \iint \pi_{nh}(dx)\overline{Q}_{nh}(x, dy)f(x, y) &= \pi_{nh} \otimes Q_{nh} \overline{f} \\ &= \pi \otimes Q \overline{f} + n^{-1/2} \pi \otimes Q h \cdot A \overline{f} + o(n^{-1/2}) \\ &= \pi \otimes Q f + n^{-1/2} \pi \otimes Q Bh \cdot f + o(n^{-1/2}) \\ &= \pi \otimes Q f + n^{-1/2} \pi \otimes Q h \cdot Af + o(n^{-1/2}). \end{aligned}$$

Combining with (3.1), we obtain the perturbation expansion

$$n^{1/2}(\pi_{nh} \otimes Q_{nh}^s f - \pi \otimes Q f) \rightarrow \pi \otimes Q h \cdot Af.$$

To prove that Q_{nh}^s is Hellinger differentiable, we must show that π_{nh} has a π -density. On $L_2(\pi)$ we have

$$\pi_{nh} = \pi(I - (Q_{nh} - Q)D)^{-1}$$

with

$$D = I + \sum_{j=1}^{\infty} (Q^j - \pi);$$

see Greenwood and Wefelmeyer (1997, Lemma 2). Write $Q_{nh} - Q = n^{-1/2}Q_{h_n}$ with $Q_h(x, dy) = Q(x, dy)h(x, y)$. Then

$$\pi_{nh} f = \pi f + \sum_{k=1}^{\infty} n^{-k/2} \pi(Q_{h_n} D)^k f.$$

As in the proof of Lemma 1 we can show that

$$\pi Q_{h_n} D f = \pi \otimes Q h_n \cdot R D f = \pi \overline{V} h_n \cdot f.$$

By induction,

$$\pi(Q_{h_n} D)^k f = \pi W^k h_n \cdot f \quad \text{for } k = 1, 2, \dots$$

with

$$W h = \overline{V} h, \quad W^k h = \overline{V}(L W^{k-1} h \cdot h) \quad \text{for } k = 2, 3, \dots$$

By Assumption 2, the operator \overline{V} is bounded on $L_2(\pi)$. We obtain a representation for the π -density of π_{nh} ,

$$d\pi_{nh}/d\pi = 1 + \sum_{k=1}^{\infty} n^{-k/2} W^k h_n \quad (5.2)$$

$$= 1 + n^{-1/2}(\overline{V} h_n + s_n) \quad (5.3)$$

with

$$s_n = \sum_{k=2}^{\infty} n^{-(k-1)/2} W^k h_n.$$

It is straightforward to check that $\pi s_n^2 \rightarrow 0$ and that $\bar{V}h_n$ and s_n are bounded by $O(n^{1/8})$.

From (5.1) and (5.2),

$$\begin{aligned} \pi_{nh}(dx)Q_{nh}(x, dy) &= \pi_{nh}(dy)\bar{Q}_{nh}(y, dx) \\ &= \pi(dy) \left(1 + n^{-1/2} \left((\bar{V}h_n)(y) + s_n(y)\right)\right) \bar{Q}_{nh}(y, dx). \end{aligned}$$

On the other hand, by definition (2.2) of Q_{nh} ,

$$\pi_{nh}(dx)Q_{nh}(x, dy) = \pi(dx) \left(1 + n^{-1/2} \left((\bar{V}h_n)(x) + s_n(x)\right)\right) Q(x, dy)(1 + n^{-1/2}h_n(x, y)).$$

Use $\pi(dx)Q(x, dy) = \pi(dy)Q(y, dx)$ and solve for \bar{Q}_{nh} to obtain that

$$\begin{aligned} \bar{Q}_{nh}(y, dx) &= Q(y, dx)(1 + n^{-1/2}h_n(x, y)) \\ &\quad \left(1 + n^{-1/2} \left((\bar{V}h_n)(x) + s_n(x)\right)\right) \\ &\quad \left(1 + n^{-1/2} \left((\bar{V}h_n)(y) + s_n(y)\right)\right)^{-1}. \end{aligned}$$

Hence \bar{Q}_{nh} is Hellinger differentiable (2.1) with derivative $\bar{h} + R\bar{V}h - L\bar{V}h$. We note that relation (3.10) implies

$$Bh = h + L\bar{V}h = \bar{h} + R\bar{V}h \quad \text{for } h \in H_0.$$

Therefore, the derivative equals h . Hence Q_{nh}^s is also Hellinger differentiable with derivative h . \square

Remark. The arguments in the proof of Lemma 3 could be used to obtain a general stability result for Markov chains of the following type: If Q' has a Q -density and is close to Q , then the invariant distribution π' of Q' has a π -density, and the density of π' has a representation in terms of the density of Q' . \square

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